

# Cross-Classified Models in the Context of Value-Added Modeling <sup>1</sup>


J. Kyle Roberts    Douglas Bates

Southern Methodist University

University of Wisconsin-Madison

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## Background to Value-Added Modeling

- Initially, VAMs focused on model development (McCaffrey, Lockwood, Koretz, Louis, & Hamilton, 2004; Tekwe et al., 2004) and the development of software routines appropriate for estimating these models.
- Work by Doran, Bates, and Phillips (under review) and work presented in the special issue of the Journal of Educational and Behavioral Statistics, volume 29(1) have made considerable strides above and beyond initial VAMs.

## Current Models for VAMs

- One of the first true VAMs was introduced by Sanders, Saxton & Horn (1997) on the Tennessee VAM System.
- VAMs typically fall into one of two categories: models appropriate for monitoring school effects on student learning and models appropriate for monitoring teacher effects on student learning.
- These models typically take the form:

$$Y_{ti} = [\beta_0 + \beta_1(\text{year}_t)] + [\theta_{0j(i)} + \theta_{1j(i)}(\text{year}_t) + \delta_{0i} + \delta_{1i}(\text{year}_t) + \epsilon_{ti}]$$

where  $i$  indexes students and  $j$  could index either the schools or teachers. The notation  $j(i)$  represents the school (or teacher) attended by student  $i$  in year  $t$ . In this model, the within-group errors are assumed IID where  $\epsilon_{ti} \sim \mathcal{N}(0, \sigma^2)$  and school (or teacher) and student random effects are  $\theta_j = (\theta_{0j}, \theta_{1j}) \sim \mathcal{N}(\mathbf{0}, \Psi)$  and  $\delta_i = (\delta_{0i}, \delta_{1i}) \sim \mathcal{N}(\mathbf{0}, \Omega)$ , respectively.

# McCaffrey et al. (2003) Recommendations for Future Research

1. Develop databases that can support VAM estimation of teacher effects.
2. Develop computational tools for fitting VAM
3. Link VAM teacher-effect estimates to alternative measures of teacher effectiveness
4. Empirically evaluate the potential sources of errors we have identified
5. Estimate the prevalence of factors that contribute to the sensitivity of teacher-effect estimates
6. Incorporate decision theory into VAM
7. Use research and auxiliary data to inform modeling choices

## The Strength of Cross-Classified VAMs

A case with true nesting.

HS	Middle School					
	MS 1	MS 2	MS 3	MS 4	MS 5	MS 6
HS 1	S1, S2	S3, S4, S5				
HS 2			S6, S7	S8, S9		
HS 3					S10, S11	S12, S13

A case with cross-classification.

HS	Middle School					
	MS 1	MS 2	MS 3	MS 4	MS 5	MS 6
HS 1	S1, S2	S3	S4		S5	
HS 2		S6, S7		S8, S9, S10		
HS 3	S11		S12		S13	S14, S15

## Notating the Cross-Classified Model

When each student is nested inside one and only one middle school and these middle schools are nested inside three high schools. This data may be modeled in a mixed-effects model as:

$$Y_{ijk} = \gamma_{000} + u_{00k} + u_{0jk} + e_{ijk}$$

where  $\gamma_{000}$  represents the average student achievement across all  $j$  middle schools and  $k$  high schools with student-level residuals  $e_{ijk}$  and middle school-level random effects  $u_{0jk}$  and high school-level random effects  $u_{00k}$ .

## Notating the Cross-Classified Model

In the cross-classified example, student scores are modeled as:

$$Y_{i(j_1, j_2)} = \gamma_{000} + u_{00j_1} + u_{00j_2} + u_{00j_1j_2} + e_{i(j_1, j_2)}$$

where  $\gamma_{000}$  represents the average student achievement for student  $i$  having attended middle school  $j_1$  and high school  $j_2$  with random effects  $u_{00j_1}$  for the middle schools,  $u_{00j_2}$  for the high schools, a random interaction effect  $u_{00j_1j_2}$  for belonging to both middle school  $j_1$  and high school  $j_2$  and student-level residuals  $e_{i(j_1, j_2)}$ . This parameterization follows the Rasbash and Browne (2001) notation.

# Rethinking the Cross-Classified VAM

- We might rethink the value-added model as

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{b} + \boldsymbol{\epsilon}, \quad \boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n), \quad \mathbf{b} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}(\boldsymbol{\theta}))$$

where  $\boldsymbol{\theta}$  is the vector of variance components to be estimated. For the model we are describing  $\boldsymbol{\theta}$  would have a total of nine elements; two variances and one covariance for each of the student, teacher and school classifications.

- Because the matrix  $\mathbf{Z}$  can be huge (perhaps millions of rows and hundreds of thousands of columns) this representation would be interesting but completely impractical if it were not for the fact that  $\mathbf{Z}$  is very, very sparse.



# Cholesky Decomposition in Sparse Matrix Theory

The broad outline of this decomposition to express the variance-covariance of the random effects,  $\Sigma(\boldsymbol{\theta})$ , in terms of a relative covariance factor,  $\Lambda(\boldsymbol{\theta})$ , defined so that

$$\Sigma(\boldsymbol{\theta}) = \sigma^2 \Lambda(\boldsymbol{\theta}) \Lambda'(\boldsymbol{\theta})$$

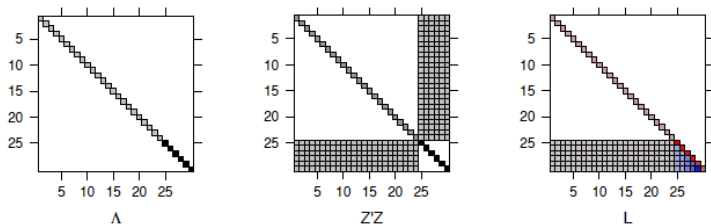
and optimize the profiled deviance, defined as

$$-2\tilde{\ell}(\boldsymbol{\theta}) = \log(|\mathbf{L}(\boldsymbol{\theta})|^2) + n \left[ 1 + \frac{2\pi r^2(\boldsymbol{\theta})}{n} \right]$$

where  $\mathbf{L}(\boldsymbol{\theta})$  is the sparse, lower-triangular Cholesky factor satisfying

$$\mathbf{L}(\boldsymbol{\theta}) \mathbf{L}'(\boldsymbol{\theta}) = \Lambda(\boldsymbol{\theta})' \mathbf{Z}' \mathbf{Z} \Lambda(\boldsymbol{\theta}) + \mathbf{I}.$$

# Graphical Representation of the Cholesky Decomposition



**Fig. 2.4** Images of the relative covariance factor,  $\Lambda$ , the cross-product of the random-effects model matrix,  $Z^T Z$ , and the sparse Cholesky factor,  $L$ , for model `fm2`.

## An Example with Only One Time Point

Our first model will be the “null” or unconditional model such that,

$$Y_{i(j_1, j_2)} = \gamma_{000} + u_{00j_1} + u_{00j_2} + e_{i(j_1, j_2)}$$

This is easily modeled in lme4 as:

```
> m0 <- lmer(ACHIEV ~ 1 + (1 | PSCHOOL) + (1 | SSCHOOL),
+           REML = F, pupils)
> summary(m0)
```

Linear mixed model fit by maximum likelihood

Formula: ACHIEV ~ 1 + (1 | PSCHOOL) + (1 | SSCHOOL)

Data: pupils

AIC	BIC	logLik	deviance	REMLdev
2326	2345	-1159	2318	2321

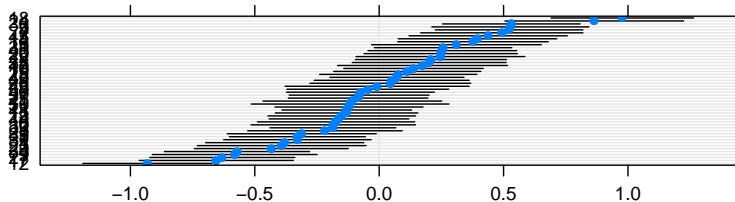
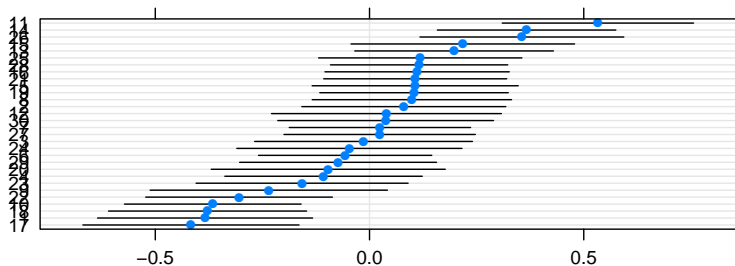
Random effects:

Groups	Name	Variance	Std.Dev.
PSCHOOL	(Intercept)	0.169348	0.41152
SSCHOOL	(Intercept)	0.065401	0.25574
Residual		0.513169	0.71636

Number of obs: 1000, groups: PSCHOOL, 50; SSCHOOL, 30

Fixed effects:

# Examining the Distributions for Secondary and Primary



## Adding Covariates

In addition to model  $m_0$ , we can test for the effect of pupil ses level and pupil gender as fixed effects such that:

$$Y_{i(j_1, j_2)} = \gamma_{000} + \gamma_{100} \text{gender}_{i(j_1, j_2)} + \gamma_{200} \text{ses}_{i(j_1, j_2)} + u_{00j_1} + u_{00j_2} + e_{i(j_1, j_2)}$$

```
> m1 <- lmer(ACHIEV ~ PUPSEX + PUPSES + (1 | PSCHOOL) +
+ (1 | SSCHOOL), REML = F, pupils)
> summary(m1)
```

Linear mixed model fit by maximum likelihood

Formula: ACHIEV ~ PUPSEX + PUPSES + (1 | PSCHOOL) + (1 | SSCHOOL)

Data: pupils

AIC	BIC	logLik	deviance	REMLdev
2255	2285	-1122	2243	2258

Random effects:

Groups	Name	Variance	Std.Dev.
PSCHOOL	(Intercept)	0.169010	0.41111
SSCHOOL	(Intercept)	0.063606	0.25220
Residual		0.474255	0.68866

Number of obs: 1000, groups: PSCHOOL, 50; SSCHOOL, 30

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	5.75548	0.10527	54.67
PUPSEXgirl	0.26131	0.04564	5.73
PUPSES	0.11409	0.01610	7.09

Correlation of Fixed Effects:

(Intr) PUPSEX

## Adding Random Effects

Furthermore, we may wish to model random effects for gender across all primary schools (and not secondary schools). This would be specified as:

$$Y_{i(j_1, j_2)} = \gamma_{000} + \gamma_{100} \text{gender}_{i(j_1, j_2)} + (\gamma_{200} + u_{10j_1}) \text{ses}_{i(j_1, j_2)} \\ + u_{00j_1} + u_{00j_2} + e_{i(j_1, j_2)}$$

```
> m2 <- lmer(ACHIEV ~ PUPSEX + PUPSES + (PUPSEX |
+ PSCHOOL) + (1 | SSCHOOL), REML = F, pupils)
> summary(m2)
```

Linear mixed model fit by maximum likelihood

Formula: ACHIEV ~ PUPSEX + PUPSES + (PUPSEX | PSCHOOL) + (1 | SSCHOOL)

Data: pupils

AIC	BIC	logLik	deviance	REMLdev
2257	2296	-1120	2241	2255

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
PSCHOOL	(Intercept)	0.144265	0.37982	
	PUPSEXgirl	0.016480	0.12838	0.463
SSCHOOL	(Intercept)	0.063742	0.25247	
Residual		0.470093	0.68563	

Number of obs: 1000, groups: PSCHOOL, 50; SSCHOOL, 30

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	5.75892	0.10278	56.03
PUPSEXgirl	0.26551	0.04900	5.42
PUPSES	0.11301	0.01608	7.03

Correlation of Fixed Effects:



# Time-Series Cross-Classified

Suppose that we have a situation where we are doing continuous progress monitoring on students three times a year across two years.

```
> xtabs(~time + cteach, vamdata)
```

	cteach									
time	10	11	12	13	14	15	20	21	23	24
0	2	2	2	2	2	2	0	0	0	0
1	2	2	2	2	2	2	0	0	0	0
2	2	2	2	2	2	2	0	0	0	0
3	0	0	0	0	0	0	3	3	3	3
4	0	0	0	0	0	0	3	3	3	3
5	0	0	0	0	0	0	3	3	3	3

```
> vam1 <- lmer(achieve ~ time + (time | cteach) +  
+ (1 | id), vamdata)  
> summary(vam1)
```

Linear mixed model fit by REML

Formula: achieve ~ time + (time | cteach) + (1 | id)

Data: vamdata

AIC	BIC	logLik	deviance	REMLdev
289.7	305.6	-137.8	276.9	275.7

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
id	(Intercept)	7.5470	2.7472	
cteach	(Intercept)	5.6589	2.3788	
	time	1.1774	1.0851	-1.000
Residual		1.0384	1.0190	

Number of obs: 72, groups: id, 12; cteach, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	2.1763	1.1172	1.948
time	1.2613	0.3535	3.568

Correlation of Fixed Effects:

(Intr)

# Random Effects for Teachers

```
> ranef(vam1)$cteacher
```

```
      (Intercept)      time
10 -2.0531202    0.9365056
11  2.1450473   -0.9784371
12 -1.9189711    0.8753152
13  0.5772548   -0.2633077
14 -2.4467006    1.1160325
15  1.4752146   -0.6729011
20  2.0513364   -0.9356920
21 -2.9085858    1.3267158
23  3.5448827   -1.6169549
24 -0.4663579    0.2127235
```